Structural Equation Modelling

Lecture 4: Bayesian Structural Equation Modelling

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Recap

- Last week: (Frequentist) SEM
- Today: Bayesian SEM $\sum_{\text{prox}_{\text{promotion}}}$

Introduction

Interest:

- Parameter estimates: loadings, residual variances, latent regressions, latent variances, latent means
- How well does the hypothesized model represent the data?

Differences to classical:

- *Fitting* the model
- Model fit is somewhat different
- Identification
- Multiple models should be considered simultaneously

Outline

- Introduction
- Parameter estimation
	- Basics
	- MCMC sampling
	- Small variance priors
- Model fit and comparison
- Multi-model inference
- Practical issues

Parameter estimation

- A parameter is a variable with a distribution that we want to approximate, the *posterior*
- The posterior distribution:
	- probability of possible parameter values after observing the data
	- Which parameter values are more likely than others
	- 95% interval contains the parameter of interest with given probability
	- What parameter value is the most likely (point estimate: mean or median)

Parameter estimation

The posterior distribution of a parameter is proportional to $p(\theta|y) \propto p(\theta) \cdot p(y|\theta)$

With:

 θ as the parameter vector

 γ as the data

 $p(\theta|y)$ as the posterior probability of the parameter given the data $p(\theta)$ as the prior probability of the parameter

 $p(y|\theta)$ as the likelihood of the data given the parameter

Prior and posterior as probability distributions

- Probability as the mass under the density curve (integral)
- Likelihood changes the prior belief towards posterior belief
- Posterior as the "compromise" between prior and likelihood

Priors in a SEM

- Prior captures belief about plausible parameter values before seeing the data
- Informative vs. non -informative?

The parameters:

: loadings

B: regressions

 Ψ : variances and covariances of the latent variables

 θ : variances and covariances of the observed residuals

 η : latent factor scores

Posterior estimation

- If we were to write the full joint posterior of all parameters: $p(A, B, \Psi, \Theta, \eta | y) \propto p(A, B, \Psi, \Theta, \eta) p(y | A, B, \Psi, \Theta, \eta)$
- An analytic solution to this becomes almost impossible
- Markov-chain Monte Carlo (MCMC) sampling allows to approximate the posterior by sampling from it
- Using software such as JAGS and Stan we could define the priors and the likelihood in the respective programming language and have the built-in MCMC sampling take care of the rest
- But: R-package *blavaan* does this
- We only need to:
	- Specify a model in *lavaan*-syntax
	- Be aware of what we are doing:
		- MCMC sampling needs to be checked
		- Parameters are no ML point estimates anymore

MCMC sampling

- Many iterations (samples) to accurately represent the target distribution
- Multiple chains to make sure if the process always leads to the same target and is independent of the starting values
- Burn-in (warm-up) to throw away early samples that are usually not very representative of the target and are more related to the starting value than to the target

Convergence diagnostics

Does the MCMC sample properly approximate the target distribution, aka, the posterior?

- R-hat:
	- Similarity of the chains
	- Should be smaller than 1.01 and close to 1.0
- Effective sample size (ESS, neff):
	- Because of autocorrelation MCMC samples not independent
	- ESS is the number of independent sample draws
	- Ideally close to the number of samples
	- *Thinning* can help: Take only every, e.g., 2nd, 4th, or 10th value of a chain
- Traceplots: visualize the chains

> blavInspect(blav_fit, what = "neff")

f_leader=~Workload 2222.686 f_leader=~Reward 1915.312 f_leader=~Fairness 1892.717 f_leader=~ValueFit 2050.289 f_resources=~LeisureBreaks 3409.622 f_resources=~WorkRelatedResources 3215.196 f_resources=~PsychosocialResources 3346.523 f_stress=~PerformanceRelatedStress 2894.546 f_burnout=~Cynicism 3030.250 HealthAwareness~~HealthAwareness 4365.089 Workload~~Workload 4230.501 Reward~~Reward 4738.278 Fairness~~Fairness 4364.824 ValueFit~~ValueFit 4869.630 OverallRecovery~~OverallRecovery 3110.338

The discovered to a substitution of a concentration of a

> blavInspect(blav_fit, what = "rhat") f_leader=~Workload 1.0004983 f_leader=~Reward 1.0008243 f_leader=~Fairness 1.0005618 f_leader=~ValueFit 1.0008286 f_resources=~LeisureBreaks 1.0005187 f_resources=~WorkRelatedResources 0.9998543 f_resources=~PsychosocialResources 0.9992834 f_stress=~PerformanceRelatedStress 1.0010401 f_burnout=~Cynicism 0.9998085 HealthAwareness~~HealthAwareness 1.0001440 Workload~~Workload 1.0003352 Reward~~Reward 1.0005263 Fairness~~Fairness 1.0000491 ValueFit~~ValueFit 0.9993236 OverallRecovery~~OverallRecovery 0.9997663 LeisureBreaks~~LeisureBreaks 0.9995578 WorkRelatedResources~~WorkRelatedResources

The traceplots look good: hairy

1000

800

600

400

200

 $1.0 -$

 \circ

200

600

800

100

400

caterpillar

 $\mathbf 0$

 $1.10 -$

1.05

ESS is good

R -hat values are within the "good"

range

Priors and sample size

- Explicit benefit of BSEM to work with small samples:
	- Large sample properties of ML are not needed: asymptotical normality
	- Convergence is not needed
	- Prior knowledge can be incorporated
- With large sample sizes, prior does rarely matter: the likelihood will dominate (the prior is "overwhelmed")
- With smaller samples it can make sense to think a bit longer about the prior, look for prior knowledge
- However: less data, less information

> mydp <- dpriors(beta="normal(1,2)") > blav_fit_pri <- bsem(model, dtFull, seed = 11, dp = mydp)

Regressions:

Results barely change

blav_fit_only <- bsem(model, dtFull, prisamp = TRUE) \swarrow $plot(blav_fit_only, pars = 1:4, plot_type = "dens")$

Samples and plots the

priors

 $plot(blav_fit, pars = 1:4, plot_type = "dens")$

Plots the posterior

Identification and small-variance priors

- Classical identification: Match between data provided information and to be estimated parameters
- In the Bayesian framework a classically unidentified model can be identified
- Instead of relying on the data to provide the information to estimate unidentified parameters, we provide the information ourselves in the form of priors
- This allows us to implement small variance priors or approximate inequality constraints
	- Improves model fit
	- Helps with modification
	- But: May conceal important misspecifications

```
> model_adjusted <- '
+ f leader =~ HealthAwareness + Workload + Reward + Fairness + ValueFit
+ f_resources =~ 0verallRecovery + LeisureBreaks + WorkRelatedResources + PsychosocialResources
+ f_stress =~ SocialEmotionalStress + PerformanceRelatedStress
+ f_burnout =~ EmotionalExhaustion + Cynicism
++ f_burnout ~ f_stress + f_resources
+ f_stress ~ f_leader + f_resources
+ f_resources ~ f_leader
                                                                           Add a small variance prior++ # small variance prior on a crossloading
+ f_leader =~ prior("normal(0, 0.08)")*OverallRecovery
+ ^{\prime}
```
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Model fit

- In general, the idea is:
	- How well does the model implied covariance matrix reproduce the data covariance matrix
- Single model fit:
	- BRMSEA, BCFI, BTLI
	- PPC: posterior predictive (model) checking
		- Posterior predictive p-value (PPP)
- Model comparison:
	- Information criteria: DIC, WAIC, LOOIC
	- Bayes factors

Model fit – the classics

• The fit function from maximum likelihood estimation

$$
F = log(|\widehat{\Sigma}|) + \text{trace}(\widehat{\Sigma}^{-1}) - \log(|\mathbf{S}|) - k
$$

can be seen as a deviance measure for the model implied covariance and the data covariance matrix

- We need the F value in the calculation of RMSEA, CFI, and TLI;
- Posterior sample of implied covariance matrices
	- -> posterior sample of F-values
	- -> posterior sample of RMSEA, CFI, and TLI: BRMSEA, BCFI, BTLI
- We also need the model complexity
- Model complexity used to be number of parameters p
- Because of the priors the model complexity is not equal to p in the Bayesian framework, it is an estimated quantity

```
> model_null <- "
```
- HealthAwareness \sim HealthAwareness
- Workload \sim Workload
- Control \sim Control
- Reward ~~ Reward
- Community ~~ Community
- Fairness \sim Fairness
- ValueFit ~~ ValueFit
- OverallRecovery ~~ OverallRecovery $\overline{+}$
- LeisureBreaks ~~ LeisureBreaks $\overline{+}$
- WorkRelatedResources ~~ WorkRelatedResources $\overline{+}$
- PsychosocialResources ~~ PsychosocialResources $\overline{+}$
- SocialEmotionalStress ~~ SocialEmotionalStress $\overline{+}$
- PerformanceRelatedStress ~~ PerformanceRelatedStress $\overline{+}$
- LossOfMeaning \sim LossOfMeaning $\overline{+}$
- EmotionalExhaustion ~~ EmotionalExhaustion $\overline{+}$
- Cynicism \sim Cynicism $\ddot{}$

```
+<sup>1</sup>
```
 $> blav_fit_null \leq bsem(model = model_null, data = dtrull)$

> blavFitIndices(blav_fit, baseline.model = blav_fit_null, fit.measures = c("BRMSEA", "BCFI", "BTLI")) Posterior mean (EAP) of devm-based fit indices:

BRMSEA BCFI BTLI 0.116 0.949 0.936 For calculation of CFI and TLI we need the null (baseline) model

We already needed that in the frequentist SEM, but lavaan did that under the hood

PPP

- Posterior predictive p-value
- Compare the discrepancy of:
	- Posterior of model implied covariance matrices and observed covariance matrix
	- Simulated data from the posterior covariance matrices and and the model implied covariance matrices (chance discrepancy, sampling distribution, a "good" fit discrepancy)
- Does the model's discrepancy differ significantly from the discrepancy expected by chance?
- Should be 0.5, the smaller the worse

> summary(blav_fit, standardized = TRUE) blavaan 0.5.3.1230 ended normally after 1000 iterations Estimator **BAYES** Optimization method **MCMC** Number of model parameters 44

491

PPP

0.000

MargLogLik

 -6878.963

Number of observations

Statistic

Value

Information criteria

- Deviance information criterion (DIC)
	- Point estimate of the deviance of the mean model implied covariance matrix
- Widely applicable information criterion (WAIC)
	- The deviance for each person for each posterior sample
- Leave-one-out cross validation information criterion (LOOIC):
	- Deviance between the model trained for N-1 data and the remaining one observation as the test data
	- N-times
	- for MCMC of SEM computationally very expensive
	- Approximation obtained with importance sampling (*blavaan* does that for you)

ELPD difference & SE: -34.466 8.453

Laplace approximation to the log-Bayes factor (experimental; positive values favor object1):

NA

Frequentist model comparison

- Compare **two** models
- H0: the models are equal
- We can only reject the HO

But

- How much more likely is the complex model than the simple model?
- What is the probability of each model given the data?
- Within the space of possible models, which is the most likely?

Bayes factor

- Comparison of two models M_1 and M_2 by comparing their marginal likelihoods (the probability of the data under the model)
- In parameter estimation: $p(y|M) = p(y)$
- The ratio of the likelihoods is the Bayes factor:

$$
BF_{12} = \frac{p(y|M_1)}{p(y|M_2)}
$$

• How much more likely are the data under M_1 than under M_2 ?

Bayes factor

- Can quantify evidence in favour of M_1 but also in favour of M_2
- For example:
	- $BF_{12} = 3$ -> the data are three times as likely under M_1 than under M_2
	- How much more likely are the data under M_2 than under M_1 ? $BF_{21} = 1/BF_{12}$
- For SEM the marginal likelihoods are not easy to estimate:
	- Approximations exist: Laplace approximation, BIC transformation for the BF
	- You could use sampling methods such as importance sampling, path sampling, bridgesampling

https://vasishth.github.io/bayescogsci/book/ch-bf.html

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Bayesian model averaging (BMA)

- With BFs we are still comparing one model against another, eventually settling on one "good" model which we base all inferences on
- What if we estimated the posterior model probability (PMP) of each possible model?
- We could use the PMP as weights for our parameter estimates

BMA

- Lets say we have 4 candidate models
- What is the posterior model probability of M1?

$$
p(M_1|y) = \frac{p(M_1) \cdot p(y|M_1)}{\sum_{j=1}^{4} p(y|M_j)p(M_j)}
$$

• In practice we often choose a uniform prior for the models, that is, all models are equally likely a-priori, so $p(M_1) = .25$

Candidate model

• Choosing candidate models can be hard if the number of possible models is large, let's only look at the regressions for now:

And many more… but let's assume we define only these four as candidate models

BMA

With the PMPs we can:

- Estimate the posterior inclusion probability of the parameter
	- Sum the posterior model probabilities that include the parameter
- Estimate a model-averaged posterior distribution for a parameter:
	- Draw a model based on the PMPs
	- Draw a value for the parameter from its posterior under this model
	- Repeat many times
	- Doing this for only the models that include the parameter answers the question, assuming the effect is present how strong is it?
- Much more: Inclusion BF, Exclusion BF

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Practical Issues for BSEM (and BMA)

- The marginal likelihood approximation is not very stable for complex models or small samples
- The priors are important for BFs and BMA
- Reducing the number of possible candidate models is not straightforward
- Comparing certain models in SEM in a BMA framework is debatable: For example, the latent variable(s) in a one-factor and two-factor model have a different meaning -> comparing \bullet and \bullet ?
- *blavaan* allows a lot of models, but not the same functionality as lavaan
- In general, lots of models do not fit well, BSEM cannot help, well calibrated test instruments and good theory are key

Good luck with the final assignment and hopefully you can enjoy the free time after $\mathbf{\mathbf{\mathbf{\mathbb{C}}}}$